

Sensorimotor adaptation affects detection of and pupil response to motor perturbations

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Abstract

Much recent research on sensorimotor adaptation has focussed on its cognitive side and how adaptation differs when it is cognitively controlled. Here, we investigate more closely what makes different types of perturbations detectable. We used different perturbation schedules to contrast how the magnitude of a motor perturbation and the associated error signal contribute to its detectability. We manipulated perturbation either following an unpredictable de Bruijn sequence, maximising the error signal, or a ramp perturbation, incrementally increasing perturbations and thus minimising error by allowing continuous adaptation. We did this in two different experiments using two perturbed everyday actions, walking with a split-belt speed perturbation (experiment 1) and grasping with a visuo-haptic size mismatch (experiment 2). In each experiment, N=48 participants were asked to judge the relevant dimension, either the relative speed of the treadmill belts, or the relative size of the objects to be grasped, while we measured participants' pupil dilation and investigated if this could serve as a no-report marker of perturbation detection. We found that participants reliably adapted to the incrementally introduced ramp perturbation and concurrently showed reduced perturbation detection relative to de Bruijn sequence perturbations. Tonic pupil-dilation parameters were sensitive to both perturbation magnitude and the type of perturbation, consistent with an effect of uncertainty. Pupil responses also differed depending on the detection response, but not sufficiently to allow trialwise classification. These patterns occurred both in walking and in grasping, with only minor differences, suggesting that the underlying mechanisms may be similar and potentially generalisable across motor actions.

Keywords: Perception & action, walking, grasping, sensorimotor adaptation, pupillometry

Running head: Perturbation detection, adaptation, and pupillometry

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1 Introduction

2 Humans are skilled at interacting with their environment. This is in large part because we learn from
3 interacting with it: Over time, our actions are continuously improved through many different processes. They
4 may become more fluent and energetically efficient (Franklin et al., 2008), more precise (Faisal et al., 2008),
5 as well as more accurate, the latter primarily through what we call sensorimotor adaptation (Krakauer &
6 Mazzoni, 2011). This can be thought of as changing a sensorimotor mapping between what is perceived
7 (often seen at a distance) and the corresponding action to minimise the error said action produces (Warren,
8 2006). Adaptation itself also comprises several processes (Tsay et al., 2024) from sensory to cognitive, and
9 can happen both implicitly, that is, without the actor deliberately changing the action, and explicitly, that is,
10 deliberately and cognitively controlled (Mazzoni & Krakauer, 2006; Smith et al., 2006).

11 Recent studies have shown major involvement of the cerebellum in controlling sensorimotor adaptation
12 (Shadmehr et al., 2010; Tseng et al., 2007), both in upper-limb (Taylor & Ivry, 2014) and lower-limb actions
13 (Hagen & Fling, 2026), with other areas like the basal ganglia, premotor cortex and primary motor cortex
14 also involved in motor skill learning (Baladron et al., 2023; Diedrichsen & Kornysheva, 2015). More recent
15 research has also investigated the role of specific brain structures with respect to individual components
16 of sensorimotor adaptation, for example with McDougle et al. (2022) finding a specific role of the medial
17 temporal lobe in acquiring memories relevant for adaptation and motor learning.

18 Sensorimotor adaptation and noticing perturbations

19 To be able to employ an explicit strategy to deal with a sensorimotor perturbation (Taylor & Ivry, 2011), the
20 actor needs to be aware of this perturbation. In experimental studies where one is interested in properties
21 of explicit adaptation, this is often manipulated (in particular in upper-limb actions) by either instructing
22 participants about the perturbation (Miyamoto et al., 2020; Taylor & Ivry, 2011), by manipulating its
23 noticeability for example by making it very large (Hudson & Landy, 2012), the outcome uniform regardless
24 of the response (McDougle et al., 2015), introducing it gradually rather than suddenly (Kagerer et al.,
25 1997; Modchalingam et al., 2019, 2023; Orban de Xivry et al., 2013), or by distracting participants from it
26 using secondary tasks (Mariscal et al., 2020), though this latter method will also show other influences on
27 adaptation patterns (Bedard & Song, 2013; Torres-Oviedo & Bastian, 2012). Indeed, already Held & Gottlieb
28 (1958) stated that it was important for experimenters to make sure that only one type of adaptation was
29 possible. But how do we as experimenters know if our method of choice achieved its goal? Early studies used
30 post-experiment interview questions to get a rough idea of whether this was achieved successfully (Gentilucci
31 et al., 1995; Säfström & Edin, 2004), but most direct tests came in the form of assessing participants’
32 “ownership” of an action (Gaffin-Cahn et al., 2019) or meta-cognitive judgements about their own performance
33 (Arbuzova et al., 2021; Pereira et al., 2023). Instead, researchers have looked at adaptation performance
34 itself (Herzfeld et al., 2014; Kagerer et al., 1997), or employed indirect no-report markers such as MEPs
35 (Orban de Xivry et al., 2013) or EMG (Albert & Shadmehr, 2016) to get an idea of whether sensing or
36 perceiving perturbations would affect adaptation to them. Until recently, few studies investigated directly
37 what perturbations to upper-limb action are perceived by participants.

38 In lower-limb actions like walking on the other hand, several studies have investigated the abilities of
39 participants to detect motor perturbations, most prominently split-belt speed perturbations. Researchers
40 typically find just-noticeable differences (JNDs) of on average around 7-9% for young and healthy participants
41 (Iturralde et al., 2020; Liss et al., 2022), but these figures can be higher for other participant groups such

42 as older adults (Lauzière et al., 2014) or patient populations (Hoogkamer et al., 2015), as well as differing
43 substantially between participants within each group (Müller & Kopiske, 2025). Similar to findings in
44 upper-limb actions, Malone & Bastian (2010) showed that distracting participants from the perturbation
45 affected the adaptation patterns. Note, however, that a similar paradigm found no effect of distraction on
46 the perception of split-belt speed differences (Liss et al., 2022), while Hoogkamer et al. (2015) reported, in
47 patients with cerebellar lesions, a correlation between the ability to detect split-belt speed differences and the
48 extent of gait adjustments made to those differences. In upper-limb actions, two studies on perturbation
49 detection in grasping from our own lab (Müller et al., 2025; Pfalz et al., 2025) found that perturbation
50 schedules that allowed participants to adapt to the perturbation correlated with worse (and decreasing)
51 detection performance, consistent with the notion that the error signal plays an important role in detecting
52 perturbations, as has been frequently suggested (Acerbi et al., 2017; Tsay et al., 2021).

53 **Pupillometry in perception, and in action**

54 In recent years, there has been increased interest in using pupil dilation as a window into cognitive and
55 perceptual processes that is unobtrusive and does not require a response from the participants. This has
56 revealed, for example, pupil-dilation signature of arousal (Nassar et al., 2012; Stanners et al., 1979), surprise
57 (Preuschhoff et al., 2011), novelty (Naber et al., 2013), changes in perceptual interpretations (Einhäuser et
58 al., 2008), and uncertainty (Lavín et al., 2014). Pertinent to the topic of perturbation detection is that it
59 has also been shown to respond to perceived physical effort (Zénon et al., 2014), multisensory integration
60 (Rigato et al., 2016), predictions about the environment (Lavín et al., 2014; Murphy et al., 2016; Vincent et
61 al., 2019), has been successfully applied in walking tasks as a marker of cognitive load (Saeedpour-Parizi et
62 al., 2020), and – likely due to it reflecting surprise (Roberts & Seidler, 2026) and uncertainty – it has been
63 shown to differ based on perturbation properties typically used to dissociate implicit and explicit adaptation
64 (Yokoi & Weiler, 2022). Thus, pupillometry can be viewed as a window into the cognitive processes involved
65 in motor tasks (Yokoi, 2025).

66 With respect to perturbation detection, pupil dilation reflecting both sensory and cognitive processes
67 opens interesting possibilities. The two major arguments against testing for perturbation detection directly
68 after each experimental trial are (i) the fact that this alerts participants to the very perturbation one may
69 not want them alerted to and (ii) typical adaptation schedules involving a repeated perturbation in the
70 same direction since this is what participants can then adapt their actions to, which in turn means that
71 participants would have to give the same answer repeatedly and can thus lead to unwanted artefacts (Bosch
72 et al., 2020). Both of these issues would be solved with a no-report marker that would eliminate the need for
73 a direct answer from the participant. To explore this possibility, we (Pfalz et al., 2025) recently tested to
74 what extent pupil dilation reflected stimulus properties in an adaptation paradigm on the one hand, and
75 response properties in a perturbation-detection task on the other hand. While certain stimulus properties
76 showed very clear effects in the pupil-dilation data, the relationship between participants’ responses and their
77 pupil dilation was only visible on average and not sufficient to allow trial-by-trial classification, which would
78 be needed for a no-report marker sufficient to replace the direct report.

79 **The present study**

80 Here we investigated three main questions. One, to what extent the reduced error signal following sensorimotor
81 adaptation leads to more difficulty detecting motor perturbations. Two, whether tonic (i.e., more long-

82 term, across trials) and phasic (i.e., more immediate) pupillary responses correlate with and are predictive
 83 of participants' abilities to detect motor perturbations. Three, if, regarding the former two aspects, the
 84 same patterns can be found in different everyday actions, namely grasping and walking, and thus if the
 85 mechanisms involved in these patterns are generalisable across actions as opposed to them being action- and
 86 effector-specific.

87 To this end, we conducted two experiments in which participants were exposed to repeated motor
 88 perturbations, introduced in such a way that adaptation to the perturbation was either facilitated or
 89 prevented (figure 1). This was done using a split-belt speed perturbation to walking in experiment 1 (figure 2),
 90 and a visuo-haptic size discrepancy in grasping in experiment 2 (figure 3), an extension of an earlier grasping
 91 study in our lab (Pfalz et al., 2025), with an improved experimental design that more clearly contrasted the
 92 respective contributions of the perturbation magnitude and the error signal. Pupil diameter was tracked
 93 throughout both experiments. We hypothesised that (i) perturbation schedules that are more easily adapted
 94 to would make it harder to detect those perturbations, (ii) pupil dilation, and especially tonic responses, as
 95 seen in our previous study (Pfalz et al., 2025), would differ between the two perturbation schedules, and (iii)
 96 that the patterns would not differ substantially between the two actions.

97 Methods

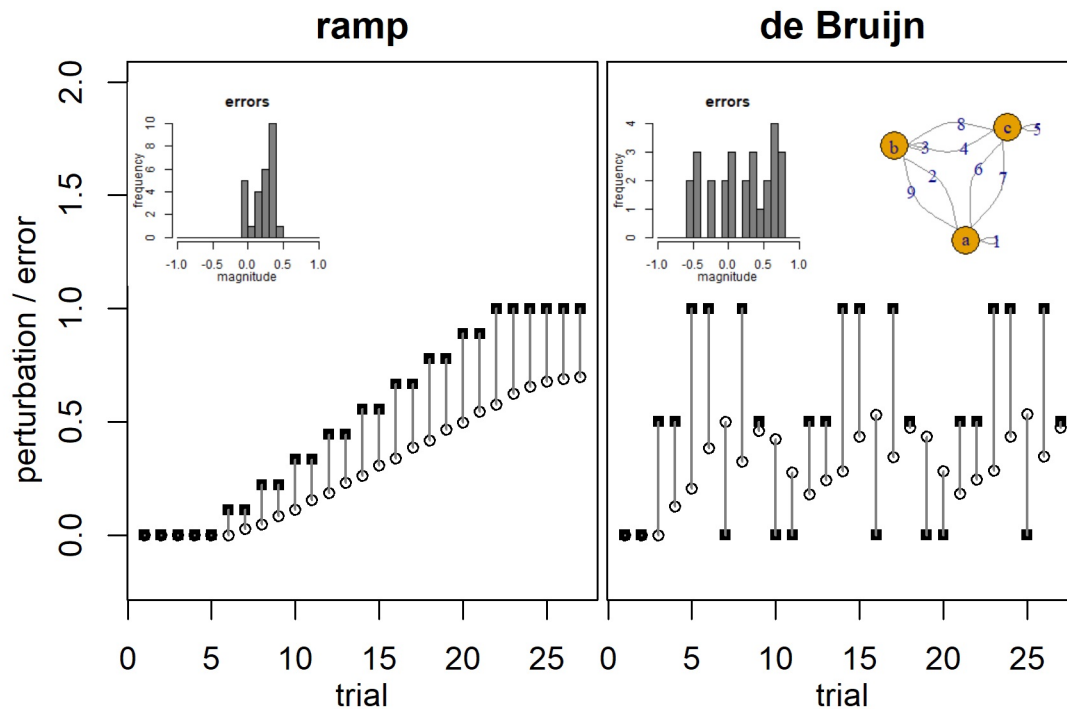


Figure 1: Different perturbation schedules and simulated errors. **Left:** Ramp-type perturbation, **right:** A thrice-repeated de Bruijn sequence for an alphabet size of three where, when looped, each element is equally often followed by every other element and by itself (top-right inset). Top-left insets show simulated errors, obtained by applying equation 1 with standard parameters of $A = 0.9$ and $b = 0.25$. The y-axes are normalised to a maximum perturbation strength of 1.

98 In both experiments, we dissociated the magnitude of the perturbation and that of the error signal. To do
99 this, we contrasted a ramp-type perturbation with the perturbation magnitude increasing in small increments
100 (used, e.g., by Mawase et al. (2013)), which allows participants to adapt to the perturbation very well, with a
101 perturbation schedule designed to provide a maximal error signal with minimal predictability. For this, we
102 employed a de Bruijn sequence (de Bruijn, 1946), which provides a minimally short sequence where, when it
103 is looped, each element (in this case: perturbation magnitude) is followed by each of the other elements and
104 itself with equal probability, which is very desirable for adaptation research (Volcic & Domini, 2018). Figure
105 1 shows the two perturbation schedules, along with simulated errors (top-left insets) using a state-space
106 model and realistic error-correction parameters taken from K. K. Kopiske et al. (2017). As the ramp-type
107 perturbation allows the participant to continuously adapt their action, there are fewer large motor errors
108 (inset).

109 Experiment 1: Perturbation detection and pupillometry in walking

110 The first experiment was conducted using a Gait Realtime Interactive Lab (GRAIL; Motek Medical, Ams-
111 terdam, NL) system, see figure 2. This system includes a dual-belt treadmill capable of accelerating and
112 decelerating each belt independently at high precision (Sessoms et al., 2014) with force plates recording
113 at 1,000 Hz under each belt, surrounded by a 240° projection screen and twelve passive motion-tracking
114 cameras (Vicon Motion Systems, Yarnton, UK) capturing the 3D positions of retro-reflective markers at
115 250 Hz. Participants also wore a Tobii 3 mobile eye-tracking device (Tobii, Stockholm, SE) that recorded
116 eye movements and pupil diameter at 100 Hz, synchronised to the data obtained by the GRAIL system via
117 custom-made triggers (Müller et al., 2023).

118 Participants

119 We aimed for a sample size of 48 to achieve a power of 0.9 with an estimated effect size of Cohen’s $d = 0.5$
120 (Cohen, 1988). We invited a total of 50 participants into the lab, of which 2 were excluded from analyses
121 due to technical difficulties during data collection. The data set used for analyses thus consisted of $N=48$
122 participants, 29 women and 19 men, with an average age of 25.6 years \pm a standard deviation of 7.2 years.
123 The average height was 173.9 ± 11.2 cm, average body mass was 69.3 ± 14.5 kg, and the average leg length
124 was 93.5 ± 6.5 cm. Body measurements were collected after the participant had completed a questionnaire
125 in which they confirmed that they had normal or corrected-to-normal vision, and felt sufficiently rested
126 and focused to complete the experiment. Participants were naïve to the hypotheses of the experiment and
127 debriefed afterwards. Having any impairment of the locomotor system was listed as an exclusion criterion in
128 the recruitment email. Compensation was either course credit or a reimbursement of 10€/h. All experimental
129 procedures were in accordance with the 2013 Declaration of Helsinki and approved by the Chemnitz University
130 of Technology ethics committee (reference no. 101628179).

131 Procedure

132 Prior to the experiment, participants first completed a questionnaire. We then took biometric measurements
133 necessary for the motion-tracking model and applied 16 retro-reflective markers following the Vicon Plug-In
134 Gait lower-body model (Vicon Motion Systems, Yarnton, UK) as well as a safety harness. This was always
135 done by the same experimenter to ensure high reliability of the measurements (McGinley et al., 2009).
136 Participants then put on the Tobii 3 eye-tracking glasses, which were also equipped with a tripod, and seven

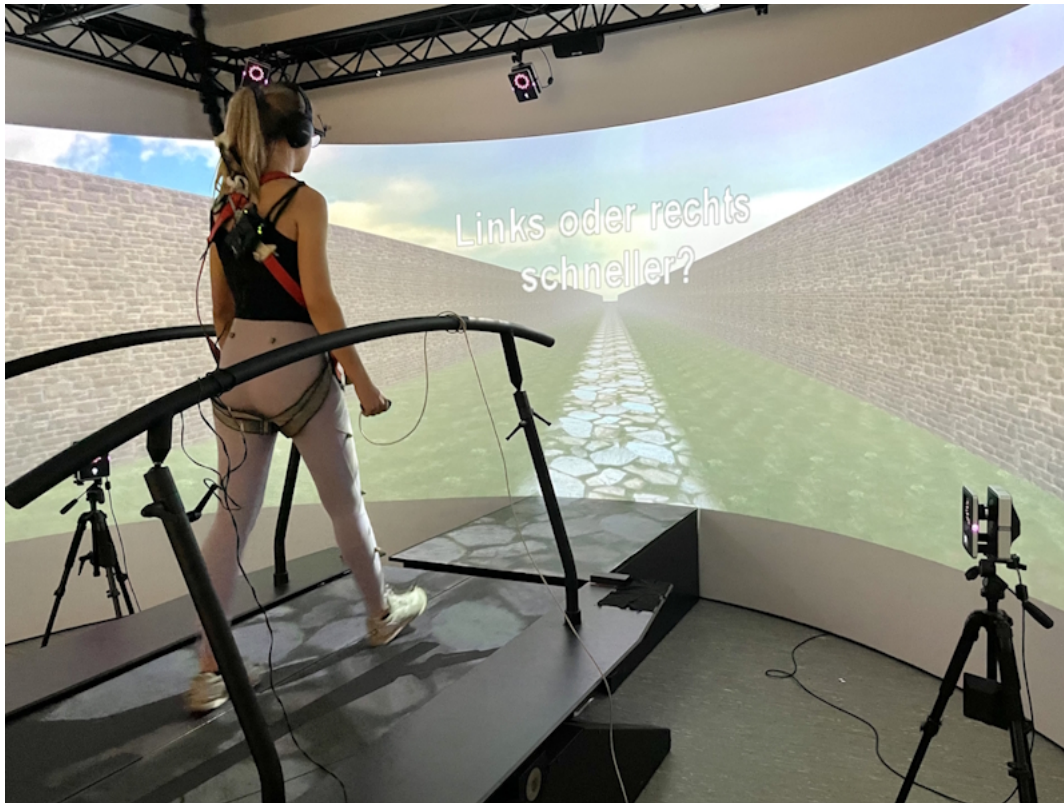


Figure 2: Setup for experiment 1. The participant is walking on a dual-belt treadmill, wearing a safety harness, eye-tracking glasses (Tobii 3 Pro), and retro-reflective motion-tracking markers. Participants also held a two-button controller in their right hand to respond to the judgement task. An endless-road scene is projected on the screen and treadmill in front of the participant. The prompt "Links oder rechts schneller?", German for "Left or right faster?", was presented three steps after each perturbation's onset.

137 additional motion-tracking markers on the head, glasses and tripod to calculate head orientation (figure 2).
 138 They then walked onto the treadmill, the experimenter attached the harness to a hook protruding from the
 139 ceiling, and the participant was ready for the experiment.

140 Each experiment started with the experimenter calibrating the motion-tracking model using an approx.
 141 ten second long recording of the participant going from a T-pose to a slow walking motion. Then, the eye
 142 tracker was calibrated using a calibration card supplied by the manufacturer. Participants then completed
 143 two short practise blocks of two minutes walking that started unperturbed at 1 m/s. After this, perturbed
 144 walking began, where each perturbation consisted of one belt changing its speed during swing time of the
 145 respective leg, followed by, after three perturbed steps, the participant being prompted to answer which side
 146 was running faster. The new speed was maintained for a minimum of 8 s, that is, until the first time the
 147 foot on the perturbed belt was lifted after 8 s had passed. If the participant did not respond within 8 s, the
 148 response was registered as not given and the participant was prompted to please respond faster. In the two
 149 practise blocks, perturbations started at 0.1 m/s speed difference between the two belts and then adaptively
 150 moved towards the participant's perceptual threshold following a QUEST algorithm (Watson & Pelli, 1983),
 151 with one practise block containing acceleration-perturbations and the other deceleration perturbations.

152 Each participant then completed a full block in which their individual JNDs for split-belt accelerations
 153 relative to 1 m/s walking were measured using a QUEST procedure starting at 0.1 m/s perturbations,

154 maximum perturbations of 0.4 m/s, and a total duration of 5 minutes (resulting in up to 41 trials). This
155 ensured that despite the large inter-individual variability in detecting such perturbations (Müller & Kopiske,
156 2025), all participants would be presented with stimuli of similar difficulty to detect. The following blocks
157 then used the final estimate of the JNDs for six further blocks: Four ramp blocks and two de Bruijn blocks.
158 In the ramp blocks, the split-belt speed difference gradually increased to either one JND difference or two
159 JNDs difference (2 x 2), with each block containing one ramp on each belt. These blocks started with 20 s of
160 unperturbed walking, followed by four ramps of which two were presented on the right belt and two on the
161 left (in randomised order) that went from 0 to the full perturbation in 7 steps. Between each ramp were 30 s
162 of unperturbed walking. In the two de Bruijn-sequence blocks, the maximal perturbation was again either
163 one or two JNDs. Blocks started with 20 s of unperturbed walking, followed by perturbations being applied
164 to one belt that followed a de Bruin sequence of 5 elements, corresponding to the maximum perturbation
165 multiplied by $\{-1, -0.5, 0, 0.5, 1\}$. The sequence was generated once offline for all participants, randomisation
166 was achieved by quasi-randomly generating, for each block, a mapping of the elements of the sequence to
167 the possible perturbation strengths, with the only constraint being that the element that came first always
168 corresponded to a perturbation of 0. In each de Bruijn-sequence block, the same sequence of perturbations
169 was applied once to one belt, followed by 16 s of unperturbed walking, and then to the other belt. This
170 ensured, as an improvement over our previous studies (Müller et al., 2025; Pfalz et al., 2025), that every
171 block included perturbations in each direction. To prevent speed changes being uncomfortably strong, we set
172 the maximum JND used in ramp and de Bruijn-sequence blocks to 0.1 m/s, meaning that participants whose
173 JND estimate from the QUEST was above 0.1 m/s were presented with the same perturbation strengths as a
174 participant with a ‘true’ measured JND of 0.1 m/s. This applied to ten out of 48 participants. Overall, the
175 experiment lasted approximately 60 minutes, excluding preparation and debriefing.

176 Data analyses

177 Data from all sources were synchronised with the help of custom triggers and custom scripts written in R
178 and run using R version 4.3.3 (R Core Team, 2024). Blink-detection was performed by the Tobii Pro Lab
179 software (Tobii, Stockholm, SE). Eye-tracking and motion-tracking data were then filtered using cubic-spline
180 interpolation and a 110-ms Savitzky-Golay filter (Savitzky & Golay, 1964) for pupil dilation and a 124-mm
181 filter for the motion-tracking data.

182 From the psychophysical responses, we computed JNDs using the R-package ‘quickpsy’ (Linares &
183 López-Moliner, 2016) and tested the notion that by comparing mean JNDs by means of t-tests and their
184 Bayesian equivalents (Morey & Rouder, 2018; Rouder et al., 2009). We also quantified the aftereffect of
185 perturbations by calculating the proportion of responses given in each direction after a perturbation ramp
186 had ended.

187 To quantify sensorimotor adaptation, we calculated, on a per-step basis, the participants’ step-length
188 asymmetry (SLA) in per cent, and peak ground-reaction-forces (GRFs), normalised to baseline by dividing
189 them by the mean peak GRFs of the first twenty steps of baseline walking. We used the parameters on the first
190 step of each perturbation as a measure of perturbation response, as well as on the first step after perturbations
191 had ended as the aftereffect. Importantly, these provide measures of both inter- and intra-leg coordination,
192 which can show different adaptation patterns (Reisman et al., 2005). To quantify the effects that both the
193 current perturbation had and the one preceding it (i.e., the adaptation aftereffect), we used the R-package
194 ‘lme4’ (Bates et al., 2015) to fit linear mixed-effect models (LMEs) in which the factors *previous perturbation*
195 *magnitude* and *perturbation schedule* were added iteratively as fixed-effects slopes to baseline models of

196 $SLA \sim perturbation_magnitude + (1|participant)$ and $GRF \sim perturbation_magnitude + (1|participant)$,
197 respectively, and discarded if they did not improve the Akaike Information Criterion, *AIC* (Akaike, 1974).

198 Trialwise pupil-dilation parameters were computed relative to the initial heel contact (iHC) of each
199 first perturbed step. Specifically, we calculated as a measure of *tonic* response the mean dilation over a
200 1,000-ms window prior to iHC, and as a measure of *phasic* response the max-min difference in the 2,500-ms
201 window following iHC (referred to as response *amplitude*). To assess how these two measures were affected
202 by perturbation properties, again we fit LMEs in which the trial properties *block number*, *trial number*,
203 blockwise *maximum perturbation*, *perturbation schedule*, *response correctness*, *time to response*, and trialwise
204 *perturbation magnitude* were added iteratively as fixed effects and discarded if they did not improve the AIC.

205 Experiment 2: Perturbation detection and pupillometry in grasping

206 Experiment 2 was a grasping task in which a mirror setup was used to dissociate the size of seen (‘visual’)
207 and grasped (‘haptic’) cuboid objects. The experiment was conducted in a motion-tracking and eye-tracking
208 lab with an Optotrak 3D-Investigator system (Northern Digital Inc., Waterloo, ON, CA) for motion tracking
209 and a table-mounted EyeLink-1000 (SR Research, Ottawa, CA) for eye tracking. Participants were seated in
210 front of a cold mirror slanted 45° away from them with their head in a chin rest, wearing occlusion goggles
211 (Milgram, 1987) and with infrared-emitting diodes for motion tracking attached to their thumb, index finger,
212 and wrist. The task was to grasp a ‘haptic’ aluminium cuboid behind the mirror while looking at another
213 ‘visual’ cuboid in front of the mirror that was projected to the same position (figure 3), and then judge the
214 relative size of the seen and grasped object, similar to our previous study (Pfalz et al., 2025).

215 Participants

216 We again aimed for a sample size of 48 to allow for a power of 0.9 with an estimated effect size of Cohen’s d
217 = 0.5, a realistic estimate as a previous study (Pfalz et al., 2025) had found an effect of Cohen’s $d = 0.69$ for
218 the main comparison of AUC differences between blocks where the perturbation was continuously changing
219 vs. when it was quickly adapted to. We invited 53 participants to the lab, of which 5 had to be replaced due
220 to technical issues: Three because of difficulties achieving acceptable eye-tracking quality, one because the
221 motion-tracking system did not record all trials, and one due to an issue with stimulus presentation. This
222 left us again with a full sample of $N=48$ participants for analyses. This sample contained 38 women and
223 10 men, with an average age of 22.4 ± 3.9 . Prior to the experiment, participants confirmed that they met
224 all inclusion criteria and felt sufficiently rested and focused to complete the experiment. They were naïve
225 to the hypotheses of the experiment and debriefed afterwards. Compensation was either course credit or
226 a reimbursement of 10€/h. All experimental procedures were in accordance with the 2013 Declaration of
227 Helsinki and approved by the Chemnitz University of Technology ethics committee (reference no. 101568507).

228 Procedure

229 Prior to the experiment, participants completed a questionnaire confirming that they had full functionality of
230 their right arm and hand, normal or correct-to-normal vision, and felt rested and alert enough to complete the
231 experiment. They then entered the experimental booth, were seated at a table in front of the experimental
232 setup, and fitted with infrared diodes on their thumb, index finger and wrist, as well as LCD goggles. After
233 being instructed about the task, participants completed a five-point eye-tracking calibration using a cardboard
234 sheet projected into the mirror. Then followed twelve practise trials that matched the experimental trials and



Figure 3: Setup for experiment 2. Participants had their head in a head rest and wore LCD goggles to prevent vision between trials. On any given trials, participants saw one of the objects on the rotating disc in the mirror and grasped an object behind the mirror that was in the same location the visual object was projected to. The eye tracker registered gaze position and pupil dilation through the cold mirror. Grasping movements were registered through diodes on the thumb, index finger and wrist, as well as one next to the haptic object that was placed such that it was visible to the 3D-Investigator when the reflective side of the object was next to it, but not visible when the object was moved, thus serving as a marker of object touch.

235 consisted of: (i) Starting with the right hand in a starting position with thumb and index finger gripping a
 236 small bump on the table to the right of the participant and the LCD goggles closed, (ii) the goggles opening
 237 and the object in the mirror in front of the participant completing a small back-and-forth movement to
 238 indicate the start of the trial, (iii) the participant performing a reach-to-grasp precision grip with thumb
 239 and index finger (Jeannerod, 1986; Napier, 1956) towards the object behind the mirror and lifting it briefly,
 240 (iv) the participant judging whether the haptic object felt *certainly smaller*, *uncertainly smaller*, *uncertainly*
 241 *larger*, or *certainly larger* than the visual one via a four-button response box ('Black Box', The Black Box
 242 ToolKit Ltd., Sheffield, UK), and (v) the goggles closing 4 s after the button press. Gaze position was
 243 tracked throughout, grasp trajectories were recorded for 5 s after opening of the goggles, which was sufficient
 244 to perform a full grasping movement even for participants who grasped very slowly. In the practise block,
 245 the experimenter asked participants which response they had intended to give to verify that the mapping
 246 of the buttons was correctly understood, and reiterated the mapping if there was any indication that the
 247 participant had misunderstood or was unsure. The experimenter also informed the participant when a 12-mm
 248 perturbation was experienced (which happened three times during practise) that this was the maximum
 249 amount by which visual and haptic object would differ.

250 Then, the main experiment began, which consisted of three blocks each (with maximal perturbations of 3
 251 mm, 6 mm, or 12 mm, chosen to range from barely noticeable to clearly noticeable following result from
 252 Hillis et al. (2002) and Müller et al. (2025)) in which perturbations followed either a ramp, or a de Bruijn
 253 sequence. Ramp blocks consisted of two unperturbed trials, followed by 14 trials in which on every other trial

254 the visuo-haptic mismatch was increased by $\frac{1}{6}$ of the maximum perturbation in the respective block. This
 255 was then followed by four unperturbed trials, and another 14 perturbed trials where the mismatch had the
 256 same magnitude as before, but the opposite sign (the order of signs was randomised), followed by another
 257 two unperturbed trials. In total, each ramp block thus consisted of 36 trials. de Bruijn-sequence blocks
 258 used a pre-generated de Bruijn sequence with five elements, which was 25 trials long. As in experiment 1,
 259 each of the five elements was quasi-randomly mapped to the block’s maximum perturbation times one of
 260 $\{-1, -0.5, 0, 0.5, 1\}$, with the only constraint that the first element of the sequence was always mapped to “no
 261 perturbation”. Overall, participants thus completed six experimental blocks and 183 experimental trials. On
 262 average, the experiment lasted around 75 minutes, including breaks but excluding preparation and debriefing.

263 Data analyses

264 Psychophysical data were modeled using the ‘quickpsy’ package (Linares & López-Moliner, 2016), obtaining
 265 JNDs that were then compared between perturbation types using t-tests and their Bayesian equivalent (Morey
 266 & Rouder, 2018; Rouder et al., 2009). We also calculated areas under the curve (AUCs) using all four
 267 responses (Green & Swets, 1966), by considering them essentially four criteria of different strictness (Naber
 268 et al., 2013). That is, each curve’s y-values (‘true positives’) were computed from physically larger objects,
 269 using the proportion of responses “definitely larger” as y1, definitely or probably larger for y2, and so on,
 270 with the x-values (‘false positives’) being computed in the same way but for smaller objects and starting with
 271 “definitely smaller”. These, too, were compared between the two conditions.

272 Motion-tracking data were filtered using cubic-spline interpolation and a 22-ms Savitzky-Golay filter
 273 (Savitzky & Golay, 1964). The start of each grasping movement was defined as the point in time when both
 274 thumb and index finger moved at more than 25 cm/s. The end point was defined as the time point when the
 275 diode next to the haptic object was no longer visible to the motion-capture system since its signal was not
 276 longer reflected by the reflective side of the object, or when it moved at more than 25 cm/s. Between these
 277 two points, the maximum euclidean distance between thumb and index finger was taken as the maximum grip
 278 aperture (MGA), our main dependent variable for grasping. To then model adaptation, we used a state-space
 279 error-correction model (Wolpert et al., 1995) in which a state, roughly corresponding to a visuomotor mapping
 280 between seen object size and the corresponding grip opening, is adapted trial by trial according to equation 1:

$$x_{t+1} = A * x_t - b * E_t \quad (1)$$

281 This was fit using the ‘Nloptr’ package in R (Ypma, 2014). We defined E_t as the difference between
 282 the observed MGA and the MGA expected given the size of the haptic object and the MGA’s response
 283 function. In other words, we defined it was the haptic error signal. The main variable of interest was the
 284 error-correction parameter b .

285 Pupil-dilation data were read into R using the EyelinkReader package (Pastukhov, 2025) and filtered
 286 using, again, cubic-spline interpolation and a 35 ms Savitzky-Golay filter (Savitzky & Golay, 1964). As
 287 measures of tonic pupil response, we calculated the average of the first 1,000 ms of each trial. To measure
 288 phasic response, we computed the min-max difference 2,500 ms post object touch (*amplitude*) – see figure 3
 289 of Pfalz et al. (2025) for a visual explanation of these parameters. These were then analysed with LMEs
 290 using the ‘lme4’ package (Bates et al., 2015) to investigate whether pupil-dilation responses depended on
 291 trial properties. To do this, predictors *perturbation magnitude*, *perturbation schedule*, *trial number*, *block*
 292 *number*, *response correctness* and *response confidence* were added iteratively and kept as a predictor if the

293 more complex model had a lower AIC (Akaike, 1974) than the simpler one. Finally, support-vector machine
 294 (SVM; Boser et al. (1992)) classifiers were used to see if trial properties, pupil-dilation trajectories, as well
 295 as pupil-dilation-velocity trajectories could be used to predict response confidence. We chose confidence
 296 as a response parameter participants had conscious access to and that would depend systematically on
 297 surprise (Yokoi & Weiler, 2022). We used the ‘e1071’ package in R (Meyer et al., 2024) for these analyses.
 298 Per participant, we used four experimental blocks as the training set and two remaining blocks – chosen
 299 quasi-randomly such that they had to differ both in perturbation magnitude and perturbation schedule –
 300 as the test set. We used a sampling frequency of 50 Hz, with equal class weights and a C parameter of 1,
 301 building on our results in Pfalz et al. (2025).

302 Results

303 Eye-tracking, psychophysical, and motion capture data from both experiments, as well as analysis scripts are
 304 available via the Open Science Framework: [https://osf.io/s4bvr/overview?view_only=3421f9c8febf4c729cec](https://osf.io/s4bvr/overview?view_only=3421f9c8febf4c729cec923a780a2882)
 305 [923a780a2882](https://osf.io/s4bvr/overview?view_only=3421f9c8febf4c729cec923a780a2882). These scripts and data allow readers to reproduce all results and figures in this manuscript.

306 Experiment 1

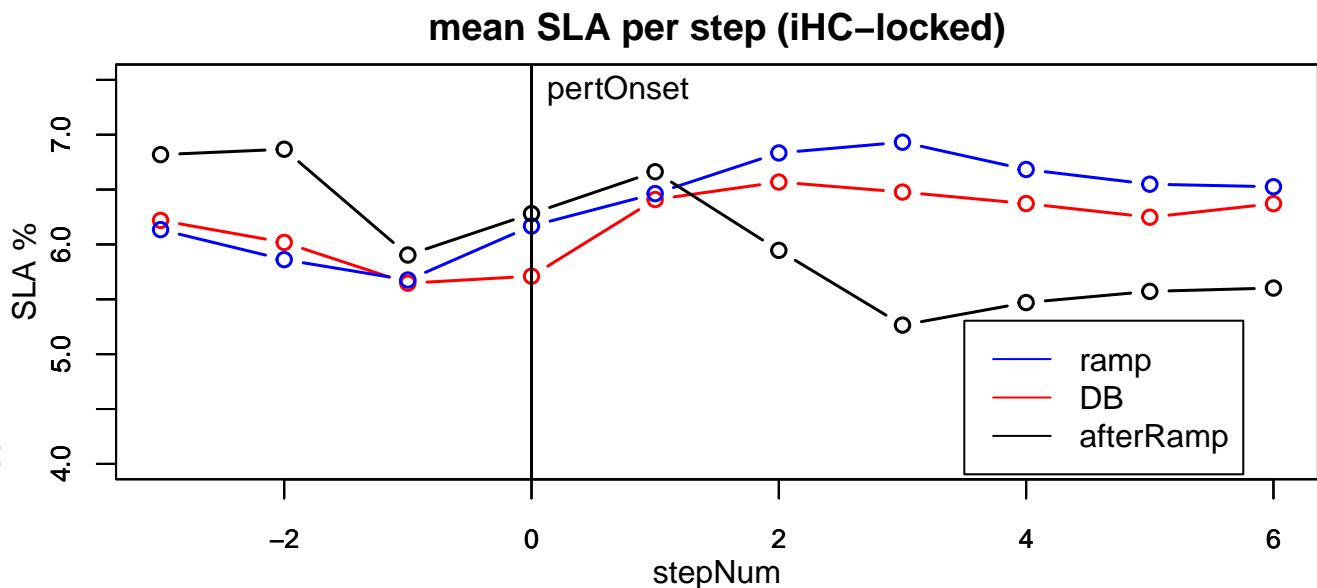


Figure 4: Step-length asymmetry per step, relative to perturbation onset. Data points show overall means across all participants and perturbation strengths, split up by block type. Initial heel contact (iHC) for the first perturbed step is used as $t = 0$. y-axis shows asymmetry as the percentage by which the longer step in any given stride was longer than the shorter step.

307 Average measured JNDs were $7.3 \% \pm$ a standard deviation of 4.1% . The minimum measured JND was
 308 1.6% , with a maximum of 27.5% and total of 10 participants’ measured thresholds were above 10% , which
 309 we had set as the maximum JND to be used in the main part of the experiment, thus setting the range to be
 310 $[1.6 \% ; 10 \%]$. Analysing data quality showed on average 0.2% of frames missing from our motion-tracking
 311 data for markers included in analyses, with only two blocks across all 48 participants exceeding 5% missing

312 data for one of those markers. Eye-tracking data were missing on average 10.7 % of frames, including blinks,
 313 ranging from 0.5 % to 34.5 % per participant.

314 Motor responses

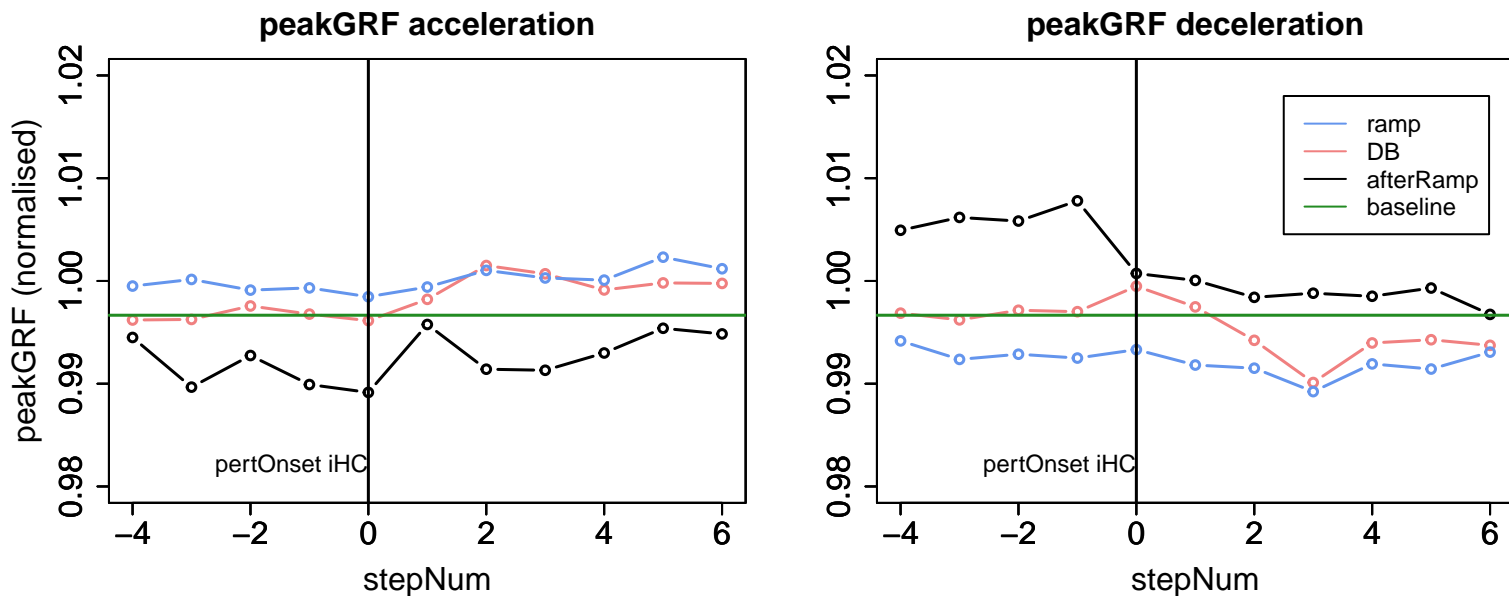


Figure 5: Normalised peak ground-reaction forces per step, relative to perturbation onset. Peaks were calculated from filtered trial-wise force-plate data. Data points show overall means for across all $N=48$ participants and perturbation strengths, split up by block type as well as accelerations (**left**) and decelerations (**right**). Forces (y-axis) were normalised to the first 20 steps of each participant’s baseline walking as 1.

315 Participants’ gait patterns showed an effect of the perturbations in increased SLA following perturbation
 316 onset, as well as following perturbation offset in ramp blocks (figure 4). LMEs show that both GRFs and SLAs
 317 were related positively to perturbation speed but negatively to preceding-perturbation speed. The full best-
 318 fitting models were $SLA \sim 21.52 * perturbation_magnitude - 21.27 * previous_perturbation_magnitude -$
 319 $1.96 * perturbation_schedule$, indicating that SLA increased by 21 % for every m/s of perturbation speed
 320 and decreased by almost the same amount for every m/s of preceding-perturbation speed and was smaller in
 321 de Bruijn blocks, and $GRF \sim 5.83 * perturbation_magnitude - 1.62 * previous_perturbation_magnitude$,
 322 showing that normalised peak GRFs were also sensitive to both current and preceding perturbation speed,
 323 but less so than SLAs, and did not depend on the block type. Similarly, peak GRFs increased following belt
 324 accelerations but decreased following decelerations in both perturbation schedules, with the opposite being
 325 true for aftereffects in ramp-type blocks (figure 5).

326 Psychophysics

327 As expected, participants showed poorer performance in the discrimination when the perturbation followed a
 328 ramp-type schedule vs. when it followed a de Bruijn sequence (figure 6). On average, participants’ JND in
 329 ramp-type perturbations were $18.9 \% \pm$ a standard error of 3.7% , compared to $12.4 \% \pm 1.8 \%$ for de Bruijn
 330 sequence blocks. This was a statistically significant difference, $t(47) = 2.32$, $p = .024$, $d = 0.34$ – however, a
 331 Bayesian t-test revealed the evidence for a difference to be only anecdotal ($BF_{10} = 1.8$). As expected, both

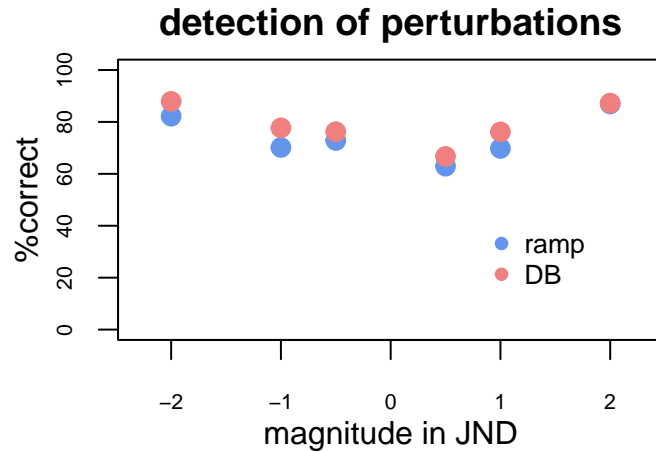


Figure 6: Percent correct in ramp and DB-sequence perturbations. Data points show overall means for all N=48 participants, split up by perturbation type and magnitude. Only perturbation magnitudes presented in both perturbation types are shown.

332 were larger than JNDs measured using the QUEST procedure, where most stimuli were near threshold. There
 333 was also the expected perceptual aftereffect where during unperturbed walking following a ramp perturbation,
 334 the belt that had been *slower* during the perturbation was, when perturbation was removed, judged as *faster*
 335 in 59.4 % of responses after 1-JND ramps, and in 78.1 % of responses after 2-JND ramps.

336 Pupil responses

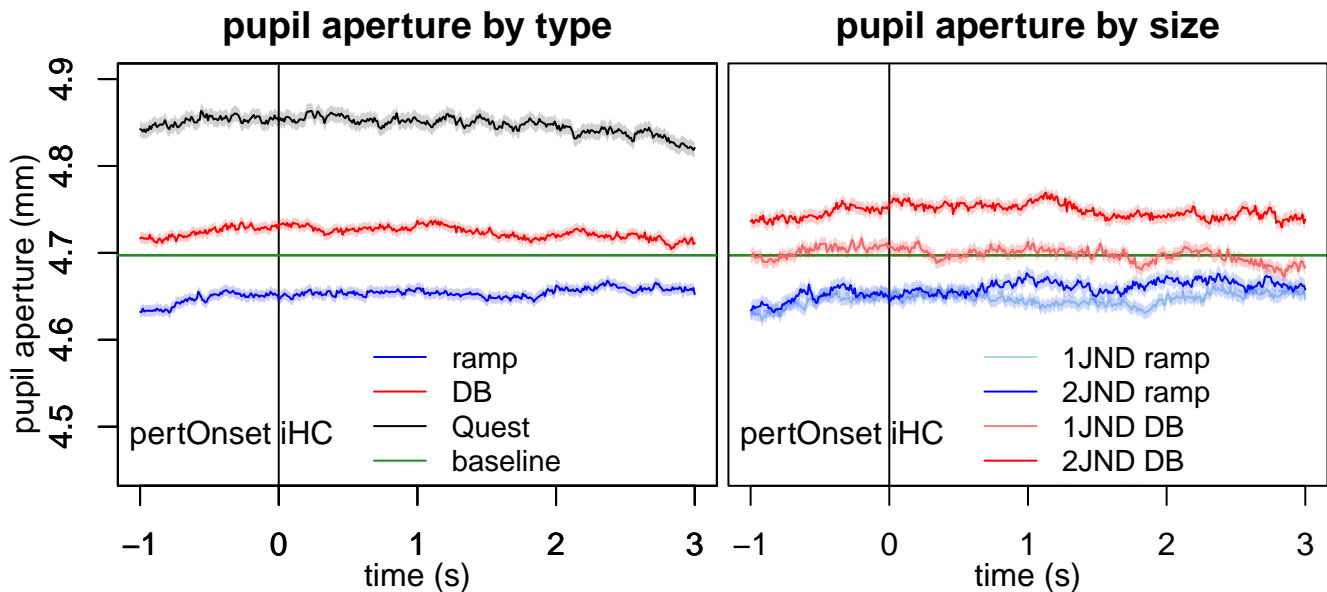


Figure 7: Mean pupil aperture relative to initial heel contact of first perturbed step. Data split up by perturbation schedule (**left**) and magnitude (**right**). We show mean collapsed trajectories over all N=48 participants, time-locked to perturbation onset. Shaded areas display standard errors, calculated as the frame-wise standard error of participant-averaged trajectories.

337 Pupil-dilation data showed a difference between trials in ramp-blocks vs. de Bruijn-sequence blocks with
 338 respect to the *tonic* pupil response. In particular, baseline dilation was higher in de Bruijn-perturbation blocks,
 339 but only by .04 mm. on average (figure 7). The full best-fitting model for pupil-dilation baseline was $baseline \sim$
 340 $-0.02 * block_number + 0.003 * trial + 0.008 * maximum_perturbation + 0.037 * perturbation_schedule +$
 341 $0.414 * perturbation_magnitude$. Neither response correctness nor time to response improved the model fit
 342 as predictors. Our phasic measure *dilation amplitude* decreased as block number increased, with the best
 343 fitting model being $amplitude \sim -0.2 * block_number$, but was not sensitive to any of the other trial or
 344 response properties, with all $\Delta AIC > 0$.

345 Experiment 1 discussion

346 In summary, results from experiment 1 showed reduced detection performance for ramp-type perturbations,
 347 but not very strongly. Both GRFs and SLAs increased when split-belt speed differences were introduced and
 348 decreased when they were removed, showing sensorimotor adaptation. Together, these finding support the
 349 notion that sensorimotor adaptation makes it harder to detect the perturbation being adapted to. Phasic
 350 pupil dilation did not differ depending on any of our predictors, whereas tonic pupil dilation differed depending
 351 on trial properties, showing increased dilation in de Bruijn-sequence blocks and when perturbations were
 352 stronger, but was independent of response properties. This shows that like in upper-limb actions (Pfalz et
 353 al., 2025; Yokoi & Weiler, 2022), pupil-dilation responses during perturbed walking differed depending on
 354 perturbation properties.

355 Experiment 2

356 Out of a total of 8,734 trials, a total of 453 trials (5.6 %) had to be excluded from motion-tracking analysis
 357 due to reaching one of the exclusion criteria laid out in the methods section. In of the remaining trials,
 358 grip-aperture data was missing for an average of 2.2 % of frames during grasping. The eye-tracking data
 359 contained more missing values, with an average of 44.7 % of frames missing, including blinks. This number
 360 was relatively high in part due to the setup with eye-tracking being recorded through occlusion goggles and a
 361 cold mirror as well as while the participant was performing an action task that required more movement
 362 than is typical for stationary eye tracking, and is in line with the data quality in our previous study (Pfalz
 363 et al., 2025). Consequently, we include analyses of data quality when reporting classification results from
 364 eye-tracking data.

365 Motor responses

366 MGA scaled with visual object size with an average slope of $.41 \pm .04$. Participants adapted their grip to the
 367 visuo-haptic mismatch with an mean error-correction parameter of $b = 0.25$. Average adaptation patterns for
 368 ramp blocks can be seen in figure 8: Especially for 6-mm and 12-mm perturbations (displayed in green and
 369 blue), there is a clear pattern of MGAs following the perturbation, as well as showing aftereffects directly
 370 after the perturbation was removed. The random assignment of perturbation sizes to the elements of the
 371 de Bruijn sequence make averaging across blocks meaningless, but analysing error-correction parameters
 372 as described in equation 1 show that adaptation was actually stronger in de Bruijn blocks (average $b =$
 373 $.30$ compared to $b = .20$ for ramp blocks). This is borne out also in the inference statistics, as we saw a
 374 main effect of *perturbation schedule* on *bs*, $F(1, 47) = 14.46$, $p < .001$, $\eta_p^2 = .24$. Neither the main effect of
 375 *perturbation magnitude* ($F(2, 94) = 1.88$, $p = .159$, $\eta_p^2 = .04$) nor the interaction between the two factors ($F(2,$

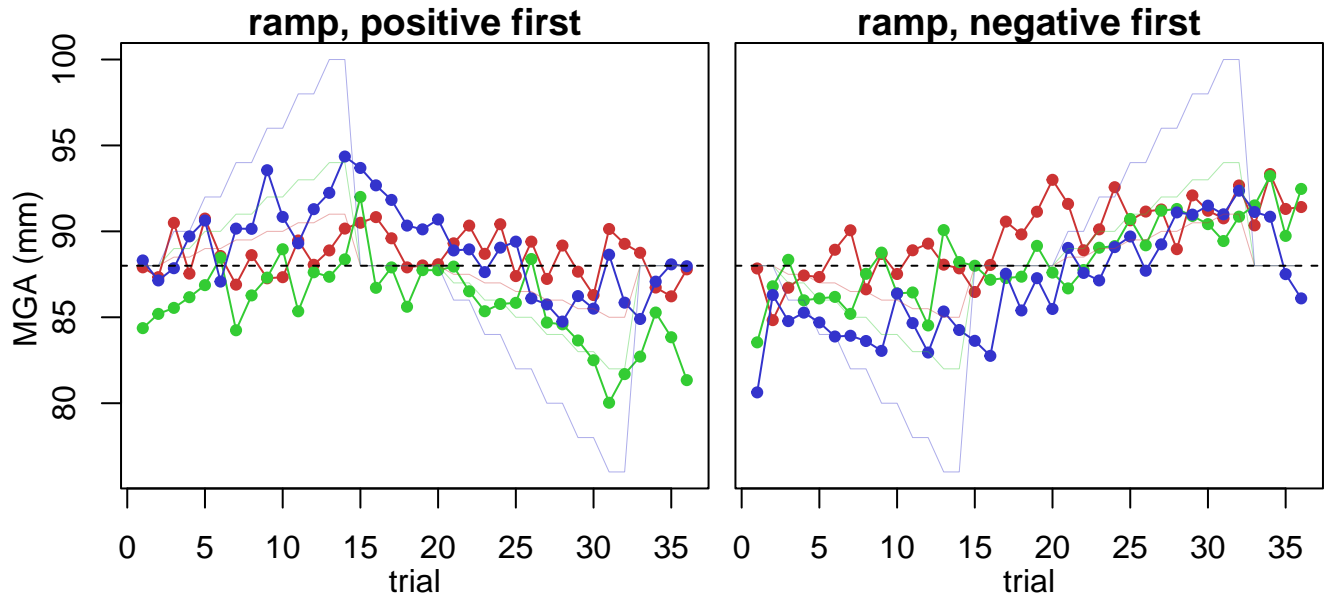


Figure 8: Grip apertures over trials in ramp blocks. Thin lines show perturbations of different strengths by trial, dots of corresponding colours show mean MGAs across all 48 participants for each perturbation. **Left:** Blocks with ramps with positive amplitude (i.e., larger haptic objects) first, followed by negative-amplitude ramps. **Right:** Negative-amplitude blocks. Data points show averages across all participants. Colours differentiate blockwise perturbation amplitudes.

376 94) = 0.04, $p = .958$, $\eta_p^2 < .01$) were statistically significant.

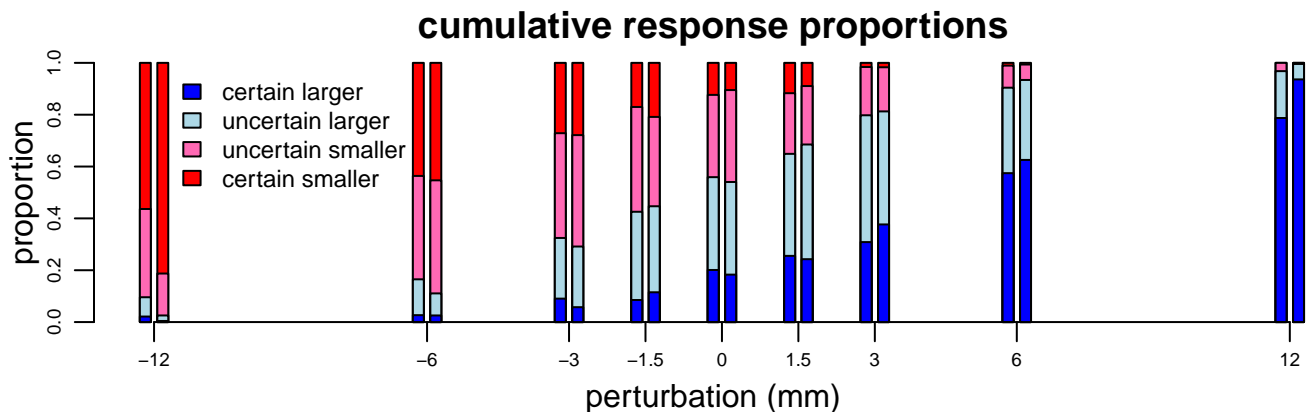


Figure 9: Cumulative response proportions by perturbation size. Perturbation size shown on the x-axis, proportions averaged across all 47 participants included in psychophysical analyses on the y-axis. Bars on the left of each pair show proportions in ramp blocks, bars on the right proportions in de Bruijn-sequence blocks. Colours differentiate response types. Data from all magnitudes present in both perturbation types are shown.

377 Psychophysics

378 One participant had to be retroactively excluded from psychophysical analyses since their larger/smaller
 379 responses did not scale with the physical differences, leaving us with a sample of N=47 for psychophysical

380 analyses. Mean JNDs were $4.71 \text{ mm} \pm 0.42$, similar to the results in our previous studies (Müller et al., 2025;
 381 Pfalz et al., 2025). The difference in JNDs between ramp blocks (5.62 mm) and de Bruijn blocks (3.81 mm)
 382 was statistically significant ($t(46) = 2.49$, $p = .016$, $d = 0.36$), although Bayesian analysis found the evidence
 383 not to be conclusive ($BF_{10} = 2.5$). The same pattern emerged when comparing AUCs (0.86 vs. 0.89; $t(46) =$
 384 -2.37 , $p = .022$, and $d = -0.35$, but $BF_{10} = 2.0$). Figure 9 shows that for all perturbation strengths, responses
 385 in ramp blocks (left bar in each pair) were more evenly distributed – that is, more similar to responses to
 386 null-perturbations – than for the same strength in de Bruijn blocks, but not by much. As in experiment 1, we
 387 also found a perceptual aftereffect of ramps, with unperturbed trials directly following ramps being judged as
 388 perturbed in the opposite direction of the preceding ramp in 59.9 % of trials following 3-mm ramps, 71.8 %
 389 of trials following 6-mm ramps, and 77.9 % of trials following 12-mm ramps.

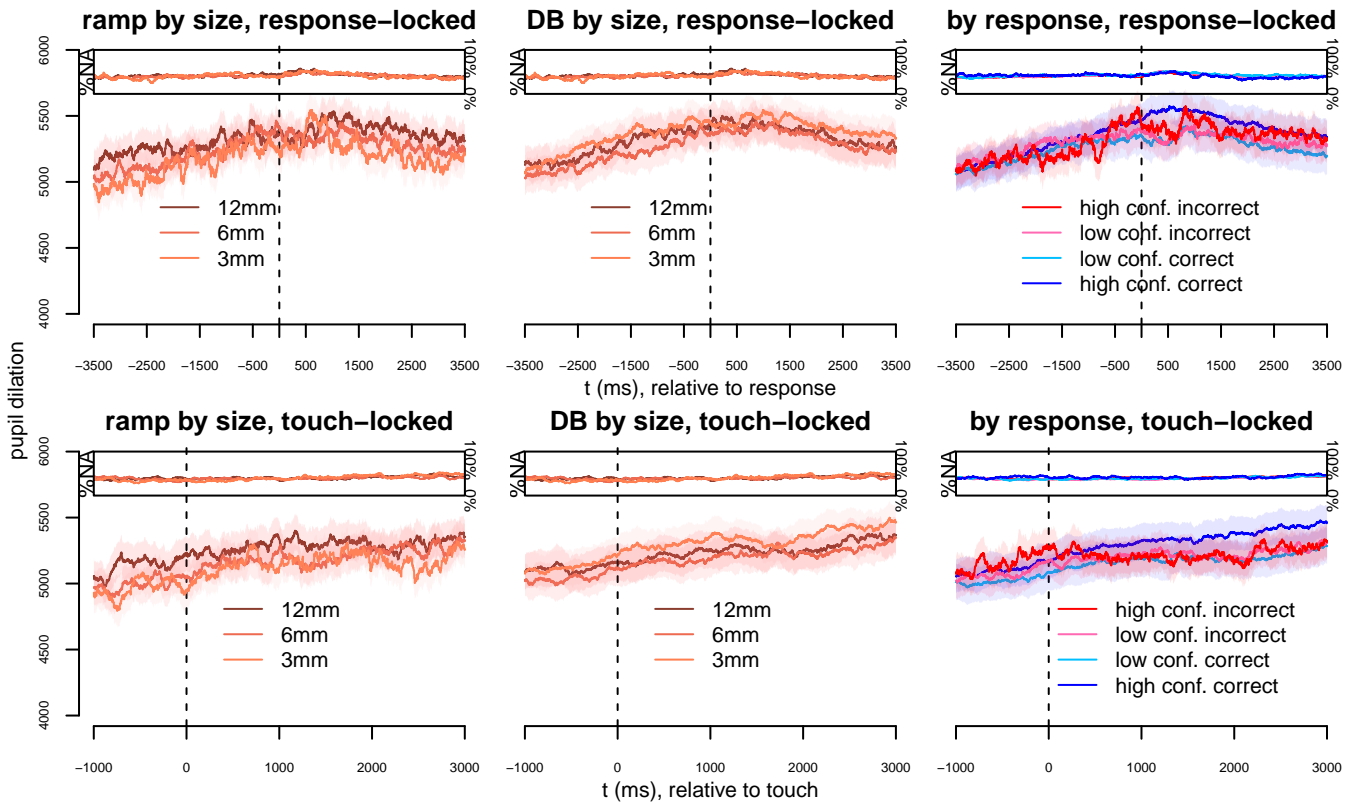


Figure 10: Pupil apertures within trials, experiment 2. Data points show trajectories split up by perturbation type and magnitude (**left** and **middle** column), as well as by response (**right** column). Trajectories are shown relative to the psychophysical response (**top** row) as well as to object touch (**bottom** row). Insets above the trajectories show the proportion of missing pupil-dilation values for each time point. Means across all trajectories of all 48 participants are shown, with shaded areas showing between-participant standard errors.

390 Pupil responses

391 With respect to pupil-dilation differences (see figure 10), LMEs revealed that the tonic response differed
 392 depending on all our predictors, with the final best-fitting model being $baseline \sim -1.03 * magnitude - 47.97 *$
 393 $perturbation_schedule - 7.84 * trial_number - 14.87 * block_number - 10.36 * response_correctness -$
 394 $9.12 * response_confidence$. For the phasic response, quantified as dilation amplitude, the best-fitting model

Table 1: SVM classification results. We show class-balanced accuracies for different predictor combinations, averaged across all N=48 participants. Average number of usable samples differed slightly depending on data quality and randomisation.

predictors	Balanced.Accuracy	Balanced.Accuracy.test	trainSamples	testSamples
trial info	0.851	0.683	49.857	25.036
eye parameters	0.725	0.479	47.346	22.731
eye samples	0.741	0.491	43.875	23.125
eye derivatives	0.962	0.518	40.727	23.455
eye samples & derivatives	0.921	0.530	43.077	22.077
info & samples & derivaties	0.932	0.485	44.737	21.789
grasping E	0.695	0.468	42.955	20.364
info + grasping E	0.893	0.721	49.429	23.179

395 $amplitude \sim 10.29 * magnitude + 2.69 * perturbation_schedule - 0.3 * trial_number - 4.67 * block_number +$
 396 $91.52 * response_correctness + 0.39 * response_confidence$ revealed that this was affected by the same
 397 predictors, but largely in the opposite direction: While dilation *baseline* was smaller for trials with stronger
 398 perturbations, in de Bruijn-schedule blocks, and with more uncertainty about the response, dilation *amplitude*
 399 was larger in all of these cases. Both parameters decreased over time. Figure 10 also illustrates, as we
 400 had found in Pfalz et al. (2025), that response-locked trajectories (top row) showed clearer patterns than
 401 touch-locked trajectories (bottom row).

402 Trialwise classification of response confidence using trial-information, pupil-dilation, and grasping error as
 403 predictors revealed that while pupil dilation and psychophysical responses correlate on average, no trial-by-trial
 404 classification in the test set was possible (table 1). Stimulus- and trial-properties were the only predictor
 405 that allowed consistently above-chance classification in the test set. As we had found previously (Pfalz et al.,
 406 2025), while pupil-dilation trajectories allowed high accuracies in the training set, including them actually
 407 *decreased* test-set accuracy compared to using only trial information. On average, trials misclassified in the
 408 test set by SVM models including pupil data had 56.7 % frames missing, compared to 55.1 % for trials that
 409 were correctly classified, showing that data quality did not explain the poor performance in the test set. Note
 410 that these numbers differ from average percentage of data missing due to resampling and different inclusion
 411 criteria for the SVM analyses.

412 Experiment 2 discussion

413 In summary, results from experiment 2 showed, just like experiment 1, that when participants were able
 414 to adapt to a perturbation (i.e., in a ramp-type perturbation schedule), their detection performance was
 415 worse than when they could not (i.e., in a de Bruijn-sequence schedule), but this was again not a large effect.
 416 Pupil-dilation parameters correlated with trial and response properties in a way that was very consistent with
 417 an effect of uncertainty, with reduced tonic and increased phasic parameters in trials with more uncertainty
 418 and surprise. For the tonic response, this is also consistent with the results from experiment 1. In fact, even in
 419 magnitude the responses were rather comparable: While the EyeLink-1000 provides pupil dilation in arbitrary
 420 units that are not directly convertible to mm (which we used in experiment 1), perturbation-schedule effects
 421 in both experiment were quite close to 1 % of total dilation. Trial-wise classification of responses was not
 422 possible using pupil-dilation data. While trial-by-trial effects were not reliable, like in Pfalz et al. (2025),
 423 these results suggest that average effects of perturbations and perturbation detection are quite consistent.

424 General discussion

425 In both experiments, participants performed more poorly on the perturbation-detection task when the
426 perturbation was introduced gradually following a ramp-type schedule, rather than an unpredictable de Bruijn
427 sequence. Thus, we replicate earlier findings from our group in grasping (Müller et al., 2025; Pfalz et al.,
428 2025) that showed that participants use both perturbation magnitude and the motor error for perturbation
429 detection, and extend them by finding the same pattern in walking, and using perturbation schedules that
430 more directly isolated the effect of a reduced error signal through adaptation. Participants corrected error
431 similarly in both perturbation schedules, consistent with other works (Coltman et al., 2021), as well as in both
432 motor tasks. Analysing pupil dilation showed consistent tonic responses to the schedule and magnitude of
433 the perturbation across both experiments, consistent with an effect of uncertainty. Phasic responses could be
434 seen in the grasping task, but not in walking. As in our previous experiment (Pfalz et al., 2025), trial-by-trial
435 classification of responses using pupil-dilation data and SVM was not successful.

436 Pupil dilation responds to perturbations, but is not a reliable no-report marker

437 As in our previous work (Pfalz et al., 2025), we found that pupil-dilation responses reflected several trial
438 properties. In particular, baseline dilation, a *tonic* response measure, was increased when perturbations were
439 stronger and more sudden. This is consistent with the notion that pupil dilation, and in particular tonic
440 responses reflect primarily the participant’s uncertainty (Dayan & Yu, 2006; Nassar et al., 2012). The fact
441 that response characteristics did not affect pupil dilation in the same way on the other hand is in line with
442 recent findings that uncertainty and errors, while correlated, so not affect motor learning in the same way
443 (Hewitson et al., 2023). *Phasic* responses on the other hand were only clearly visible in grasping, and even
444 there did not show a coherent picture. We note, however, that seeing no phasic response in walking may
445 not necessarily show that there was none, but may also point to methodological issues. For example, in our
446 previous study (Pfalz et al., 2025), we showed marked difference between touch-locked and response-locked
447 pupil trajectories, emphasising the importance of analysing relative to a suitable time point. While iHC
448 seemed to us as *a priori* the most sensible starting time point to analyse pupil-dilation trajectories during
449 perturbed walking, it could be that this was not an ideal choice. The time of the psychophysical responses
450 was typically so late and so variable between trials and participants that we did not consider it sensible.
451 Furthermore, unlike in grasping, walking provided continuously more information as the participant took
452 more perturbed steps, making it a more prudent strategy to accumulate evidence before giving a more patient
453 response. That said, from this it does not follow that iHC was the optimal choice. Also in line with our
454 previous work (Pfalz et al., 2025), trialwise classification of responses through pupil data was not possible,
455 and while data quality was not ideal, we found again that percentage of missing data was virtually identical
456 in trials classified correctly and incorrectly. Indeed, including pupil-dilation features *reduced* classification
457 performance in the test set relative to using only trial-information features, which in combination with the
458 very good training-set classification corroborates our conclusion from Pfalz et al. (2025) that this is likely a
459 consequence of overfitting.

460 These findings have implications both for our understanding of how actors adapt under perceptual
461 uncertainty, and methodological implications for using pupillometry as a no-report marker of perturbation
462 detection. On the former, we not only replicate the pattern that error signal and perturbation magnitude
463 each contribute to perturbation detection, we also show that the clearest physiological responses were tonic
464 responses that are typically associated with uncertainty or, as Yokoi (2025) wrote (p. 5), participants being

465 aware of environmental change. Consistently, these parameters were more predictive of behaviour than
466 surprise-related phasic pupil responses, which underlines that when it comes to dealing with potential motor
467 errors, the key is to make actions robust to them, since to be able to act under difficult conditions, ‘prevention
468 is better than cure’ (Warren et al., 1986) (p. 259). In walking, markers of such behaviour are multiple, as
469 participants not only adapt directly to perturbations, but also show altered gaze (K. Kopiske et al., 2021)
470 and gait patterns (Dingwell & Cusumano, 2000; Rand et al., 1998) when walking is perturbed. Future
471 research may investigate if such markers can also be seen in grasping, although we note that the cost of not
472 successfully grasping an object is of course not comparable to the cost of falling during perturbed walking.
473 For the use of pupillometry as a no-report marker, our results show that psychophysical responses at this
474 point cannot be replaced by pupil-dilation data. We have conducted three studies each with different motor
475 actions, perturbation schedules, and psychophysical responses types (Müller et al., 2025; Pfalz et al., 2025),
476 with remarkable consistency across results obtained from the different paradigms. In the two of those studies
477 that included pupillometry, we have also twice found the same pattern that pupil-dilation patterns are robust
478 on average, but very variable on a trial-by-trial basis.

479 **Similar data patterns in two very different motor actions**

480 By choosing walking and grasping as our model actions, we used two inherently multimodal actions. To
481 grasp and lift an object, an actor will use visual information about its size, position, as well as material
482 properties (Schmidt et al., 2017), information about their own hand from vision (Volcic & Domini, 2016) as
483 well as proprioception and somatosensation (van Beers et al., 1999; Weiler et al., 2019), and finally haptic
484 information when the object is touched (Lederman et al., 1986). It also forced participants to make essentially
485 cross-modal intensity comparisons, where Molyneux’s problem famously asks whether such a comparison is
486 at all possible beyond simply relying on past experiences (Locke, 1690), which we know affect perceptual
487 judgements through short-term and long-term priors (Constant et al., 2025). Similarly, walking, especially
488 under conditions of greater difficulty, will also involve vestibular (Nashner et al., 1982), but also visual
489 (Laurent & Thomson, 1988), proprioceptive or haptic cues (Dietz et al., 1987). These of course differ from
490 those used in grasping, in terms of the limbs, muscles and joints, as well as innervation involved (Boyd, 1980;
491 Nashner, 1976) as well as the time scales (Crevecoeur et al., 2020; K. Kopiske et al., 2021). The key point is
492 that choosing two experimental designs with all of these factors involved was a deliberate choice to make the
493 actions more generalisable and applicable to everyday life, where all actions have multimodal components, as
494 was already noted by Woodworth (1899) (p. 72). Finding robust effects of adaptation affecting perturbation
495 detection and correlating with tonic pupil responses *despite* the multitude of modalities and factors involved
496 underlines the generalisability of our findings.

497 **Limitations and outlook**

498 Our study improved on our previous works by making the results more generalisable, both by replicating
499 them in a different motor action, but also by using perturbation schedules that were chosen specifically to
500 isolate perturbation magnitude and the error signal. We also removed the previous issue of presenting in
501 some blocks only perturbations with the same sign, which can lead to response biases masking the perceptual
502 effects (Bosch et al., 2020). Here, all blocks contained an equal number of perturbations in each direction.
503 The key result of reduced detection performance when participants could adapt to the perturbation was
504 generalisable and robust to these methodological improvements – however, we note that it was also notably

505 reduced. This in turn may indicate that while participants do rely on their motor errors for such judgements,
506 the role of the perturbation magnitude itself is stronger than we had previously concluded. We also concede
507 that data-quality issues with our eye tracking in the grasping setup should be taken into account when
508 interpreting these data, and in particular the failure of classification which, if it worked, would be one of the
509 key advantages of establishing pupil dilation as a marker of perturbation detection. While we believe that
510 ultimately, ecologically valid experimental designs like ours are the best test of its utility as such a marker, it
511 may be interesting in future work to employ designs closer to classic psychophysical studies to see how far
512 pupillometry can go as a predictor of perception and cognition in action.

513 **Conclusion**

514 We show that adapting to a perturbation impairs the actor's ability to detect the perturbation, in walking
515 and grasping similarly. We also demonstrate concurrent tonic pupil-dilation responses that are consistent
516 with an effect of uncertainty. While pupil dilation could not be established as a reliable no-report marker
517 of perturbation detection, this sheds more light on the role of cognition and strategies in sensorimotor
518 adaptation.

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